

Justicia: A Stochastic SAT Approach to Formally Verify Fairness

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Abstract

As a technology ML is oblivious to societal good or bad, and thus, the field of fair machine learning has stepped up to propose multiple mathematical definitions, algorithms, and systems to ensure different notions of fairness in ML applications. Given the multitude of propositions, it has become imperative to formally verify the fairness metrics satisfied by different algorithms on different datasets. In this paper, we propose a *stochastic satisfiability* (SSAT) framework, Justicia, that formally verifies different fairness measures of supervised learning algorithms with respect to the underlying data distribution. We instantiate Justicia on multiple classification and bias mitigation algorithms, and datasets to verify different fairness metrics, such as disparate impact, statistical parity, and equalized odds. Justicia is scalable, accurate, and operates on non-Boolean and compound sensitive attributes unlike existing distribution-based verifiers, such as FairSquare and VeriFair. Being distribution-based by design, Justicia is more robust than the verifiers, such as AIF360, that operate on specific test samples. We also theoretically bound the finite-sample error of the verified fairness measure.

1. Introduction

Machine learning (ML) is becoming the omnipresent technology of our time. ML algorithms are being used for high-stake decisions like college admissions, crime recidivism, insurance, and loan decisions etc. Thus, human lives are now pervasively influenced by data, ML, and their inherent bias.

Example 1 *Let us consider an example (Figure 1) of deciding eligibility for health insurance depending on the fitness and income of the individuals of different age groups (20-40 and 40-60). Typically, incomes of individuals increase as their ages increase while their fitness deteriorate. We assume relation of income and fitness depends on the age as per the Normal distributions in Figure 1. Now, if we train a decision tree (Narodytska et al., 2018) on these fitness and income indicators to decide the eligibility of an individual to get a health insurance,*

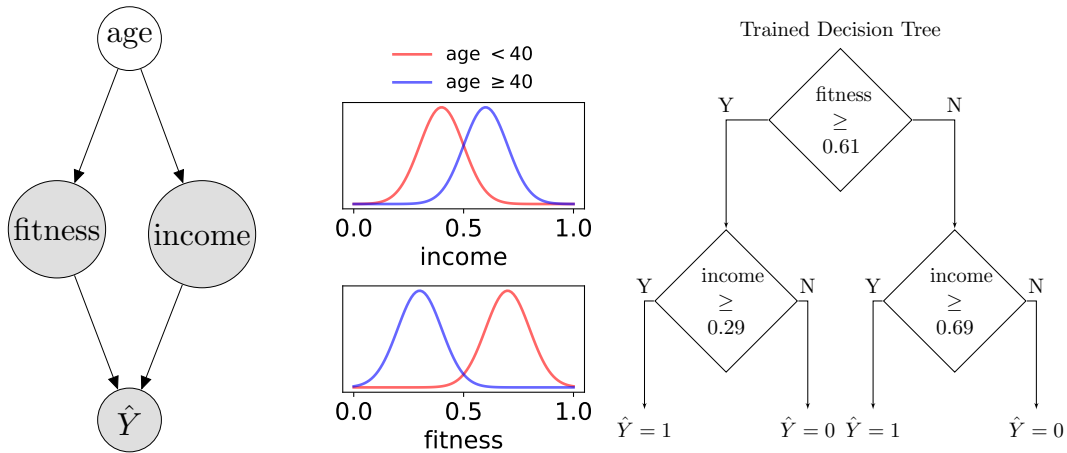


Figure 1: A trained decision tree to learn eligibility for health insurance using age-dependent fitness and income indicators.

we observe that the ‘optimal’ decision tree (ref. Figure 1) selects a person above and below 40 years with probabilities 0.18 and 0.72 respectively. This simple example demonstrates that even if an ML algorithm does not explicitly learn to differentiate on the basis of a sensitive attribute, it discriminates different age groups due to the utilitarian sense of accuracy that it tries to optimize.

Fair ML. Statistical discriminations caused by ML algorithms have motivated researchers to develop several frameworks to ensure fairness and several algorithms to mitigate bias. Existing fairness metrics mostly belong to three categories: *independence*, *separation*, and *sufficiency* (Mehrabi et al., 2019). Independence metrics, such as demographic parity, statistical parity, and group parity, try and ensure the outcomes of an algorithm to be independent of the groups that the individuals belong to (Feldman et al., 2015; Dwork et al., 2012). Separation metrics, such as equalized odds, define an algorithm to be fair if the probability of getting the same outcomes for different groups are same (Hardt et al., 2016). Sufficiency metrics, such as counterfactual fairness, constrain the probability of outcomes to be independent of individual’s sensitive data given their identical non-sensitive data (Kusner et al., 2017).

In Figure 1, independence is satisfied if the probability of getting insurance is same for both the age groups. Separation is satisfied if the number of ‘actually’ (ground-truth) ineligible and eligible people getting the insurance are same. Sufficiency is satisfied if the eligibility is independent of their age given their attributes are the same. Thus, we see that the metrics of fairness can be contradictory and complimentary depending on the application and the data (Corbett-Davies and Goel, 2018). Different algorithms have also been devised to ensure one or multiple of the fairness definitions. These algorithms try to rectify and mitigate the bias in the data and thus in the prediction-model in three ways: *pre-processing* the data (Kamiran and Calders, 2012; Zemel et al., 2013; Calmon et al., 2017), *in-processing* the algorithm (Zhang et al., 2018), and *post-processing* the outcomes (Kamiran et al., 2012; Hardt et al., 2016).

Fairness Verifiers. Due to the abundance of fairness metrics and difference in algorithms to achieve them, it has become necessary to verify different fairness metrics over datasets and algorithms.

In order to verify fairness as a model property on a dataset, verifiers like *FairSquare* (Albarghouthi et al., 2017) and *VeriFair* (Bastani et al., 2019) have been proposed. These verifiers are referred to as *distributional verifiers* owing to the fact that their inputs are a probability distribution of the attributes in the dataset and a model of a suitable form, and their objective is to verify fairness w.r.t. the distribution and the model. Though FairSquare and VeriFair are robust and has asymptotic convergence guarantees, we observe that they scale up poorly with the size of inputs and also do not generalize to non-Boolean and compound sensitive attributes. In contrast to the distributional verifiers, another line of work, referred to as sample-based verifiers, has focused on the design of testing methodologies on a given fixed data sample (Galhotra et al., 2017; Bellamy et al., 2018). Since sample-based verifiers are dataset-specific, they generally do not provide robustness over the distribution.

Thus, a *unified formal framework* to verify *different fairness metrics* of an ML algorithm, which is *scalable*, capable of *handling compound protected groups*, *robust* with respect to the test data, and *operational on real-life* datasets and fairness-enhancing algorithms, is missing in the literature.

Our Contribution. From this vantage point, *we propose to model verifying different fairness metrics as a Stochastic Boolean Satisfiability (SSAT) problem* (Littman et al., 2001). SSAT was originally introduced by (Papadimitriou, 1985) to model *games against nature*. In this work, we primarily focus on reductions to the exist-random quantified fragment of SSAT, which is also known as E-MAJSAT (Littman et al., 2001). SSAT is a conceptual framework that has been employed to capture several fundamental problems in AI such as computation of maximum a posteriori (MAP) hypothesis (Fremont et al., 2017), propositional probabilistic planning (Majercik, 2007), circuit verification (Lee and Jiang, 2018) and so on. Furthermore, our choice of SSAT as a target formulation is motivated by the recent algorithmic progress that has yielded efficient SSAT tools (Lee et al., 2017, 2018).

Our contributions are summarised below:

- We propose a unified SSAT-based approach, *Justicia*, to verify independence and separation metrics of fairness for different datasets and classification algorithms.
- Unlike previously proposed formal distributional verifiers, namely FairSquare and VeriFair, *Justicia* verifies fairness for compound and non-Boolean sensitive attributes.
- Our experiments validate that our method is more accurate and scalable than the distributional verifiers, such as FairSquare and VeriFair, and more robust than the sample-based empirical verifiers, such as AIF360.
- We prove a finite-sample error bound on our estimated fairness metrics which is stronger than the existing asymptotic guarantees.

It is worth remarking that significant advances in AI bear testimony to the right choice of formulation, for example, formulation of planning as SAT (Kautz et al., 1992). In this context, we view that formulation of fairness as SSAT has potential to spur future work

from both the modeling and encoding perspective as well as core algorithmic improvements in the underlying SSAT solvers.

2. Background: Fairness and SSAT

In Section 2.1, we define different fairness metrics for a supervised learning problem. Following that, we discuss Stochastic Boolean Satisfiability (SSAT) problem in Section 2.2.

2.1 Fairness Metrics for Machine Learning

Let us represent a dataset D as a collection of triads (X, A, Y) sampled from an underlying data generating distribution \mathcal{D} . $X \triangleq \{X_1, \dots, X_m\} \in \mathbb{R}^m$ is the set of non-protected (or non-sensitive) attributes. $A \triangleq \{A_1, \dots, A_n\}$ is the set of categorical protected attributes. Y is the binary label (or class) of (X, A) . A compound protected attribute $\mathbf{a} = \{a_1, \dots, a_n\}$ is a valuation to all A_i 's and represents a *compound protected* group. For example, $A = \{\text{race}, \text{sex}\}$, where $\text{race} \in \{\text{Asian}, \text{Colour}, \text{White}\}$ and $\text{sex} \in \{\text{female}, \text{male}\}$. Thus, $\mathbf{a} = \{\text{Colour}, \text{female}\}$ is a compound protected group. We define $\mathcal{M} \triangleq \Pr(\hat{Y}|X, A)$ to be a binary classifier trained from samples in the distribution \mathcal{D} . Here, \hat{Y} is the predicted label (or class) of the corresponding data.

As we illustrated in Example 1, a classifier \mathcal{M} that solely optimizes accuracy, i.e., the average number of times $\hat{Y} = Y$, may discriminate certain compound protected groups over others (Chouldechova and Roth, 2020). Now, we describe two family of fairness metrics that compute bias induced by a classifier and are later verified by Justicia.

2.1.1 INDEPENDENCE METRICS OF FAIRNESS.

The *independence (or calibration) metrics* of fairness state that the output of the classifier should be independent of the compound protected group. A notion of independence is referred to *group fairness* that specifies an *equal positive predictive value (PPV) across all compound protected groups* for an algorithm \mathcal{M} , i.e., $\Pr[\hat{Y} = 1|A = \mathbf{a}, \mathcal{M}] = \Pr[\hat{Y} = 1|A = \mathbf{b}, \mathcal{M}], \forall \mathbf{a}, \mathbf{b} \in A$. Since satisfying group fairness exactly is hard, relaxations of group fairness, such as *disparate impact* and *statistical parity* (Dwork et al., 2012; Feldman et al., 2015), are proposed.

Disparate impact (DI) (Feldman et al., 2015) measures the ratio of PPVs between the most favored group and least favored group, and prescribe it to be close to 1. Formally, a classifier satisfies $(1 - \epsilon)$ -disparate impact if, for $\epsilon \in [0, 1]$,

$$\min_{\mathbf{a} \in A} \Pr[\hat{Y} = 1|\mathbf{a}, \mathcal{M}] \geq (1 - \epsilon) \max_{\mathbf{b} \in A} \Pr[\hat{Y} = 1|\mathbf{b}, \mathcal{M}].$$

Another popular relaxation of group fairness, *statistical parity* (SP) measures the difference of PPV among the compound groups, and prescribe this to be near zero. Formally, an algorithm satisfies ϵ -statistical parity if, for $\epsilon \in [0, 1]$,

$$\max_{\mathbf{a}, \mathbf{b} \in A} |\Pr[\hat{Y} = 1|\mathbf{a}, \mathcal{M}] - \Pr[\hat{Y} = 1|\mathbf{b}, \mathcal{M}]| \leq \epsilon.$$

For both disparate impact and statistical parity, lower value of ϵ indicates higher group fairness of the classifier \mathcal{M} .

2.1.2 SEPARATION METRICS OF FAIRNESS.

In the *separation (or classification parity)* notion of fairness, the predicted labels \hat{Y} of a classifier \mathcal{M} is independent of the sensitive attributes A given the actual class labels Y . In case of binary classifiers, a popular separation metric is *equalized odds* (EO) (Hardt et al., 2016) that computes the difference of false positive rates (FPR) and the difference of true positive rates (TPR) among all compound protected groups. Lower value of equalized odds indicates better fairness. A classifier \mathcal{M} satisfies ϵ -equalized odds if, for all compound protected groups $\mathbf{a}, \mathbf{b} \in A$,

$$\begin{aligned} |\Pr[\hat{Y} = 1|A = \mathbf{a}, Y = 0] - \Pr[\hat{Y} = 1|A = \mathbf{b}, Y = 0]| &\leq \epsilon, \\ |\Pr[\hat{Y} = 1|A = \mathbf{a}, Y = 1] - \Pr[\hat{Y} = 1|A = \mathbf{b}, Y = 1]| &\leq \epsilon. \end{aligned}$$

In this paper, we formulate verifying the aforementioned independence and separation metrics of fairness as stochastic Boolean satisfiability (SSAT) problem, which we define next.

2.2 Stochastic Boolean Satisfiability (SSAT)

Let $\mathbf{B} = \{B_1, \dots, B_m\}$ be a set of Boolean variables. A *literal* is a variable B_i or its complement $\neg B_i$. A propositional formula ϕ defined over \mathbf{B} is in *Conjunctive Normal Form (CNF)* if ϕ is a conjunction of clauses and each clause is a disjunction of literals. Let σ be an assignment to the variables $B_i \in \mathbf{B}$ such that $\sigma(B_i) \in \{1, 0\}$ where 1 is logical TRUE and 0 is logical FALSE. The propositional *satisfiability* problem (SAT) (Biere et al., 2009) finds an assignment σ to all $B_i \in \mathbf{B}$ such that the formula ϕ is evaluated to be 1. In contrast to the SAT problem, the *Stochastic Boolean Satisfiability* (SSAT) problem (Littman et al., 2001) is concerned with the probability of the satisfaction of the formula ϕ . An SSAT formula is of the form

$$\Phi = Q_1 B_1, \dots, Q_m B_m, \phi, \tag{1}$$

where $Q_i \in \{\exists, \forall, \mathfrak{P}^{p_i}\}$ is either of the existential (\exists), universal (\forall), or randomized (\mathfrak{P}^{p_i}) quantifiers over the Boolean variable B_i and ϕ is a quantifier-free CNF formula. In the SSAT formula Φ , the quantifier part $Q_1 B_1, \dots, Q_m B_m$ is known as the *prefix* of the formula ϕ . In case of randomized quantification \mathfrak{P}^{p_i} , $p_i \in [0, 1]$ is the probability of B_i being assigned to 1. Given an SSAT formula Φ , let B be the outermost variable in the prefix. The satisfying probability of Φ can be computed by the following *rules*:

1. $\Pr[\text{TRUE}] = 1, \Pr[\text{FALSE}] = 0,$
2. $\Pr[\Phi] = \max_B \{\Pr[\Phi|_B], \Pr[\Phi|_{\neg B}]\}$ if B is existentially quantified,
3. $\Pr[\Phi] = \min_B \{\Pr[\Phi|_B], \Pr[\Phi|_{\neg B}]\}$ if B is universally quantified,
4. $\Pr[\Phi] = p \Pr[\Phi|_B] + (1 - p) \Pr[\Phi|_{\neg B}]$ if B is randomized quantified with probability p of being TRUE,

where $\Phi|_B$ and $\Phi|_{\neg B}$ denote the SSAT formulas derived by eliminating the outermost quantifier of B by substituting the value of B in the formula ϕ with 1 and 0 respectively. In this paper, we focus on two specific types of SSAT formulas: *random-exist* (RE) SSAT and *exist-random* (ER) SSAT. In the ER-SSAT (resp. RE-SSAT) formula, all existentially (resp.

randomized) quantified variables are followed by randomized (resp. existentially) quantified variables in the prefix.

Lemma 1 *Solving the ER-SSAT and RE-SSAT problems are NP^{PP} hard (Littman et al., 2001).*

The problem of SSAT and its variants have been pursued by theoreticians and practitioners alike for over three decades (Majercik and Boots, 2005; Fremont et al., 2017; Huang et al., 2006). We refer the reader to (Lee et al., 2017, 2018) for detailed survey. It is worth remarking that the past decade have witnessed a significant performance improvements thanks to close integration of techniques from SAT solving with advances in weighted model counting (Sang et al., 2004; Chakraborty et al., 2013, 2014).

3. Justicia: An SSAT Framework to Verify Fairness Metrics

In this section, we present the primary contribution of this paper, *Justicia*, which is an SSAT-based framework for verifying independence and separation metrics of fairness.

Given a binary classifier \mathcal{M} and a probability distribution over dataset $(X, A, Y) \sim \mathcal{D}$, our goal is to verify whether \mathcal{M} achieves independence and separation metrics with respect to the distribution \mathcal{D} . We focus on a classifier that can be translated to a CNF formula of Boolean variables \mathbf{B} . The probability p_i of $B_i \in \mathbf{B}$ being assigned to 1 is induced by the data generating distribution \mathcal{D} . In order to verify fairness metrics in compound protected groups, we discuss an enumeration-based approach in Section 3.1 and an equivalent learning-based approach in Section 3.2. We conclude this section with a theoretical analysis for a high-probability error bound on the fairness metric in Section 3.3.

3.1 Evaluating Fairness with RE-SSAT Encoding

In order to verify independence and separation metrics, the core component of *Justicia* is to compute the positive predictive value $\Pr[\hat{Y} = 1 | A = \mathbf{a}]$ for a compound protected group \mathbf{a} . For simplicity, we initially make some assumptions and discuss their practical relaxations in Section 3.4. We first assume the classifier \mathcal{M} is representable as a CNF formula, namely $\phi_{\hat{Y}}$, such that $\hat{Y} = 1$ when $\phi_{\hat{Y}}$ is satisfied and $\hat{Y} = 0$ otherwise. Since a Boolean CNF classifier is defined over Boolean variables, we assume all attributes in X and A to be Boolean. Finally, we assume independence of non-protected attributes on protected attributes and p_i is the probability of the attribute X_i being assigned to 1 for any $X_i \in X$.

Now, we define a RE-SSAT formula $\Phi_{\mathbf{a}}$ to compute the probability $\Pr[\hat{Y} = 1 | A = \mathbf{a}]$. In the prefix of $\Phi_{\mathbf{a}}$, all non-protected Boolean attributes in X are assigned randomized quantification and they are followed by the protected Boolean attributes in A with existential quantification. The CNF formula ϕ in $\Phi_{\mathbf{a}}$ is constructed such that ϕ encodes the event inside the target probability $\Pr[\hat{Y} = 1 | A = \mathbf{a}]$. In order to encode the conditional $A = \mathbf{a}$, we take the conjunction of the Boolean variables in A that symbolically specifies the compound protected group \mathbf{a} . For example, we represent two protected attributes: $\text{race} \in \{\text{White}, \text{Colour}\}$ and $\text{sex} \in \{\text{male}, \text{female}\}$ by the Boolean variables R and S respectively. Thus, the compound groups $\{\text{White}, \text{male}\}$ and $\{\text{Colour}, \text{female}\}$ are represented by $R \wedge S$ and $\neg R \wedge \neg S$,

respectively. Thus, the RE-SSAT formula for computing the probability $\Pr[\hat{Y} = 1|A = \mathbf{a}]$ is

$$\Phi_{\mathbf{a}} := \underbrace{\forall^{p_1} X_1, \dots, \forall^{p_m} X_m}_{\text{non-protected attributes}}, \underbrace{\exists A_1, \dots, \exists A_n}_{\text{protected attributes}}, \phi_{\hat{Y}} \wedge (A = \mathbf{a}).$$

In $\Phi_{\mathbf{a}}$, the existentially quantified variables A_1, \dots, A_n are assigned values according to the constraint $A = \mathbf{a}$.¹ Therefore, by solving the SSAT formula $\Phi_{\mathbf{a}}$, the SSAT solver finds the probability $\Pr[\Phi_{\mathbf{a}}]$ for the protected group $A = \mathbf{a}$ given the random values of X_1, \dots, X_m , which is the PPV of the protected group \mathbf{a} for the distribution \mathcal{D} and algorithm \mathcal{M} .

For simplicity, we have described computing the PPV of each compound protected group without considering the correlation between the protected and non-protected attributes. In reality, correlation exists between the protected and non-protected attributes. Thus, they may have different conditional distributions for different protected groups. We incorporate these conditional distributions in `Justicia_enum` by evaluating the conditional probability $p_i = \Pr[X_i = \text{TRUE}|A = \mathbf{a}]$ instead of the independent probability $\Pr[X_i = \text{TRUE}]$ for any $X_i \in X$. We illustrate this method in Example 2.

Example 2 (RE-SSAT encoding) *Here, we illustrate the RE-SSAT formula for calculating the PPV for the protected group ‘age ≥ 40 ’ in the decision tree of Figure 1. We assign three Boolean variables F, I, J for the three nodes in the tree such that the literal F, I, J denote ‘fitness ≥ 0.61 ’, ‘income ≥ 0.29 ’, and ‘income ≥ 0.69 ’, respectively. We consider another Boolean variable A where the literal A represents the protected group ‘age ≥ 40 ’. Thus, the CNF formula for the decision tree is $(\neg F \vee I) \wedge (F \vee J)$. From the distribution in Figure 1, we get $\Pr[F] = 0.41$, $\Pr[I] = 0.93$, and $\Pr[J] = 0.09$. Given this information, we calculate the PPV for the protected group ‘age ≥ 40 ’ by solving the RE-SSAT formula:*

$$\Phi_A := \forall^{0.41} F, \forall^{0.93} I, \forall^{0.09} J, \exists A, (\neg F \vee I) \wedge (F \vee J) \wedge A.$$

From the solution to this SSAT formula, we get $\Pr[\Phi_A] = 0.43$. Similarly, to calculate the PPV for the group ‘age < 40 ’, we replace the unit (single-literal) clause A with $\neg A$ in the CNF in Φ_A and construct another SSAT formula $\Phi_{\neg A}$ where $\Pr[\Phi_{\neg A}] = 0.43$. Therefore, if $\Pr[F], \Pr[I], \Pr[J]$ are computed independently of A and $\neg A$, both age groups demonstrate equal PPV as the protected attribute is not explicitly present in the classifier. However, there is an implicit bias in the data distribution for different protected groups and the classifier unintentionally learns it. To capture this implicit bias, we calculate the conditional probabilities $\Pr[F|A] = 0.01$, $\Pr[I|A] = 0.99$, and $\Pr[J|A] = 0.18$ from the distribution. Using the conditional probabilities in Φ_A , we find that $\Pr[\Phi_A] = 0.18$ for ‘age ≥ 40 ’. For ‘age < 40 ’, we similarly obtain $\Pr[F|\neg A] = 0.82$, $\Pr[I|\neg A] = 0.88$, and $\Pr[J|\neg A] = 0.01$, and thus $\Pr[\Phi_{\neg A}] = 0.72$. Thus, `Justicia_enum` detects the discrimination of the classifier among different protected groups. An astute reader would observe that I and J are not independent. Following (Chavira and Darwiche, 2008), we can simply capture relationship between the variables using constraints and if needed, auxiliary variables. In this case, it suffices to add the constraint $J \rightarrow I$.

1. An RE-SSAT formula becomes an R-SSAT formula when the assignment to the existential variables are fixed.

Algorithm 1 Justicia: SSAT-based Fairness Verifier

```

1: function Justicia_enum( $X, A, \hat{Y}$ )
2:    $\phi_{\hat{Y}} := \text{CNF}(\hat{Y} = 1)$ 
3:   for all  $\mathbf{a} \in A$  do
4:      $p_i \leftarrow \text{CalculateProb}(X_i|\mathbf{a}), \forall X_i \in X$ 
5:      $\phi := \phi_{\hat{Y}} \wedge (A = \mathbf{a})$ 
6:      $\Phi_{\mathbf{a}} := \forall^{p_1} X_1, \dots, \forall^{p_m} X_m, \exists A_1, \dots, \exists A_n, \phi$ 
7:      $\text{Pr}[\Phi_{\mathbf{a}}] \leftarrow \text{SSAT}(\Phi_{\mathbf{a}})$ 
8:   return  $\max_{\mathbf{a}} \text{Pr}[\Phi_{\mathbf{a}}], \min_{\mathbf{a}} \text{Pr}[\Phi_{\mathbf{a}}]$ 
9: function Justicia_learn( $X, A, \hat{Y}$ )
10:   $\phi_{\hat{Y}} := \text{CNF}(\hat{Y} = 1)$ 
11:   $p_i \leftarrow \text{CalculateProb}(X_i), \forall X_i \in X$ 
12:   $\Phi_{\text{ER}} := \exists A_1, \dots, \exists A_n, \forall^{p_1} X_1, \dots, \forall^{p_m} X_m, \phi_{\hat{Y}}$ 
13:   $\Phi'_{\text{ER}} := \exists A_1, \dots, \exists A_n, \forall^{p_1} X_1, \dots, \forall^{p_m} X_m, \neg \phi_{\hat{Y}}$ 
14:  return  $\text{SSAT}(\Phi_{\text{ER}}), 1 - \text{SSAT}(\Phi'_{\text{ER}})$ 

```

Measuring Fairness Metrics. As we compute the probability $\text{Pr}[\hat{Y} = 1|A = \mathbf{a}]$ by solving the SSAT formula $\Phi_{\mathbf{a}}$, we use $\text{Pr}[\Phi_{\mathbf{a}}]$ to measure different fairness metrics. For that, we compute $\text{Pr}[\Phi_{\mathbf{a}}]$ for all compound groups $\mathbf{a} \in A$ that requires solving exponential (with n) number of SSAT instances. We elaborate this enumeration approach, `Justicia_enum`, in Algorithm 1 (Line 1–8).

We calculate the ratio of the minimum and the maximum probabilities according to the definition of disparate impact in Section 2. We compute statistical parity by taking the difference between the maximum and the minimum probabilities of all $\text{Pr}[\Phi_{\mathbf{a}}]$. Moreover, to measure equalized odds, we compute two SSAT instances for each compound group with modified values of p_i . Specifically, to compute TPR, we use the conditional probability $p_i = \text{Pr}[X_i|Y = 1]$ on samples with class label $Y = 1$ and take the difference between the maximum and the minimum probabilities of all compound groups. In addition, to compute FPR, we use the conditional probability $p_i = \text{Pr}[X_i|Y = 0]$ on samples with $Y = 0$ and take the difference similarly. Thus, `Justicia_enum` allows us to compute different fairness metrics using a unified algorithmic framework.

3.2 Learning Fairness with ER-SSAT Encoding

In most practical problems, there can be exponentially many compound groups based on the different combinations of valuation to the protected attributes. Therefore, the enumeration approach in Section 3.1 may suffer from scalability issues. Hence, we propose efficient SSAT encodings to *learn* the most favored group and the least favored group for given \mathcal{M} and \mathcal{D} , and to compute their PPVs to measure different fairness metrics.

Learning the Most Favored Group. In an SSAT formula Φ , the order of quantification of the Boolean variables in the prefix carries distinct interpretation of the satisfying probability of Φ . In ER-SSAT formula, the probability of satisfying Φ is the *maximum* satisfying probability over the existentially quantified variables given the randomized quantified

variables (by Rule 2, Sec. 2.2). In this paper, we leverage this property to compute the most favored group with the highest PPV. We consider the following ER-SSAT formula.

$$\Phi_{\text{ER}} := \exists A_1, \dots, \exists A_n, \forall^{p_1} X_1, \dots, \forall^{p_m} X_m, \phi_{\hat{Y}}. \quad (2)$$

The CNF formula $\phi_{\hat{Y}}$ is the CNF translation of the classifier $\hat{Y} = 1$ without any specification of the compound protected group. Therefore, as we solve Φ_{ER} , we find the assignment to the existentially quantified variables $A_1 = a_1^{\max}, \dots, A_n = a_n^{\max}$ for which the satisfying probability $\Pr[\Phi_{\text{ER}}]$ is maximum. Thus, we compute the most favored group $\mathbf{a}_{\text{fav}} \triangleq \{a_1^{\max}, \dots, a_n^{\max}\}$ achieving the highest PPV.

Learning the Least Favored Group. In order to learn the least favored group in terms of PPV, we compute the *minimum* satisfying probability of the classifier $\phi_{\hat{Y}}$ given the random values of the non-protected variables X_1, \dots, X_m . In order to do so, we have to solve a ‘universal-random’ (UR) SSAT formula (Eq. (3)) with universal quantification over the protected variables and randomized quantification over the non-protected variables (by Rule 3, Sec. 2.2).

$$\Phi_{\text{UR}} := \forall A_1, \dots, \forall A_n, \forall^{p_1} X_1, \dots, \forall^{p_m} X_m, \phi_{\hat{Y}}. \quad (3)$$

A UR-SSAT formula returns the minimum satisfying probability of ϕ over the universally quantified variables in contrast to the ER-SSAT formula that returns the maximum satisfying probability over the existentially quantified variables. Due to practical issues to solve UR-SSAT formula, in this paper, we leverage the *duality* between UR-SSAT (Eq. (3)) and ER-SSAT formulas (Eq. (4))

$$\Phi'_{\text{ER}} := \exists A_1, \dots, \exists A_n, \forall^{p_1} X_1, \dots, \forall^{p_m} X_m, \neg\phi_{\hat{Y}}. \quad (4)$$

and solve the UR-SSAT formula on the CNF ϕ using the ER-SSAT formula on the complemented CNF $\neg\phi$ (Littman et al., 2001). Lemma 2 encodes this duality.

Lemma 2 *Given Eq. (3) and (4), $\Pr[\Phi_{\text{UR}}] = 1 - \Pr[\Phi'_{\text{ER}}]$.*

As we solve Φ'_{ER} , we obtain the assignment to the protected attributes $\mathbf{a}_{\text{unfav}} \triangleq \{a_1^{\min}, \dots, a_n^{\min}\}$ that maximizes Φ'_{ER} . If p is the maximum satisfying probability of Φ'_{ER} , according to Lemma 2, $1 - p$ is the minimum satisfying probability of Φ_{UR} , which is the PPV of the least favored group $\mathbf{a}_{\text{unfav}}$. We present the algorithm for this learning approach, namely `Justicia_learn` in Algorithm 1 (Line 9–14).

In ER-SSAT formula of Eq. (4), we need to negate the classifier $\phi_{\hat{Y}}$ to another CNF formula $\neg\phi_{\hat{Y}}$. The naïve approach of negating a CNF to another CNF generates exponential number of new clauses. Here, we can apply Tseitin transformation that increases the clauses linearly while introducing linear number of new variables (Tseitin, 1983). As an alternative, we also directly encode the classifier \mathcal{M} for the negative class label $\hat{Y} = 0$ as a CNF formula and pass it to Φ'_{ER} , if possible. The last approach is generally more efficient than the other approaches as the resulting CNF is often smaller.

Example 3 (ER-SSAT encoding) *Here, we illustrate the ER-SSAT encodings for learning the most favored and the least favored group in presence of multiple protected groups. As*

the example in Figure 1 is degenerate for this purpose, we introduce another protected group ‘sex $\in \{\text{male, female}\}$ ’. Consider a Boolean variable S for ‘sex’ where the literal S denotes ‘sex = male’. With this new protected attribute, let the classifier be $\mathcal{M} \triangleq (\neg F \vee I \vee S) \wedge (F \vee J)$, where F, I, J have same distributions as discussed in Example 2. Hence, we obtain the ER-SSAT formula of \mathcal{M} to learn the most favored group:

$$\Phi_{\text{ER}} := \exists S, \exists A, \mathcal{Y}^{0.41} F, \mathcal{Y}^{0.93} I, \mathcal{Y}^{0.09} J, (\neg F \vee I \vee S) \wedge (F \vee J).$$

As we solve Φ_{ER} , we learn that the assignment to the existential variables $\sigma(S) = 1, \sigma(A) = 0$, i.e. ‘male individuals with age < 40 ’ is the most favored group with PPV computed as $\Pr[\Phi_{\text{ER}}] = 0.46$. Similarly, to learn the least favored group, we negate the CNF of the classifier \mathcal{M} to obtain the following ER-SSAT formula:

$$\Phi_{\text{ER}'} := \exists S, \exists A, \mathcal{Y}^{0.41} F, \mathcal{Y}^{0.93} I, \mathcal{Y}^{0.09} J, \neg((\neg F \vee I \vee S) \wedge (F \vee J)).$$

Solving $\Phi_{\text{ER}'}$, we learn the assignment $\sigma(S) = 0, \sigma(A) = 0$ and $\Pr[\Phi_{\text{ER}'}] = 0.57$. Thus, ‘female individuals with age < 40 ’ constitute the least favored group with PPV: $1 - 0.57 = 0.43$. Thus, Justicia_learn allows us to learn the most and least favored groups and the corresponding discrimination.

We use the PPVs of the most and least favored groups to compute fairness metrics as described in Section 3.1. We prove equivalence of Justicia_enum and Justicia_learn in Lemma 3.

Lemma 3 *Let $\Phi_{\mathbf{a}}$ be the RE-SSAT formula for computing the PPV of the compound protected group $\mathbf{a} \in A$. If Φ_{ER} is the ER-SSAT formula for learning the most favored group and Φ_{UR} is the UR-SSAT formula for learning the least favored group, then $\max_{\mathbf{a}} \Pr[\Phi_{\mathbf{a}}] = \Pr[\Phi_{\text{ER}}]$ and $\min_{\mathbf{a}} \Pr[\Phi_{\mathbf{a}}] = \Pr[\Phi_{\text{UR}}]$.*

3.3 Theoretical Analysis: Error Bounds

We access the data generating distribution through finite number of samples observed from it. These finite sample set introduce errors in the computed probabilities of the randomised quantifiers being 1. These finite-sample errors in computed probabilities induce further errors in the computed positive predictive value (PPV) and fairness metrics. In this section, we provide a bound on this finite-sample error.

Let us consider that \hat{p}_i is the estimated probability of a Boolean variable B_i being assigned to 1 from k -samples and p_i is the true probability according to \mathcal{D} . Thus, the true satisfying probability p of Φ is the weighted sum of all satisfying assignments of the CNF ϕ : $p = \sum_{\sigma} \prod_{B_i \in \sigma} p_i$. This probability is estimated as \hat{p} using k -samples from the data generating distribution \mathcal{D} such that $\hat{p} \leq \epsilon_0 p$ for $\epsilon_0 \geq 1$.

Theorem 4 *For an ER-SSAT problem, the sample complexity is given by*

$$k = O\left((n + \ln(1/\delta)) \frac{\ln m}{\ln \epsilon_0}\right),$$

where $\frac{\hat{p}}{p} \leq \epsilon_0$ with probability $1 - \delta$ such that $\epsilon_0 \geq 1$.

Table 1: Results on synthetic benchmark. ‘—’ refers that the verifier cannot compute the metric.

Metric	Exact	Justicia	FairSquare	VeriFair	AIF360
Disparate impact	0.26	0.25	0.99	0.99	0.25
Stat. parity	0.53	0.54	—	—	0.54

Corollary 5 *If k samples are considered from the data-generating distribution in Justicia such that*

$$k = O\left((n + \ln(1/\delta))\frac{\ln m}{\ln \epsilon_0}\right),$$

the estimated disparate impact \hat{DI} and statistical parity \hat{SP} satisfy, with probability $1 - \delta$, $\hat{DI} \leq \epsilon_0 DI$, and $\hat{SP} \leq \epsilon_0 SP$.

3.4 Extending to Practical Settings

In this section, we relax assumptions of Boolean classifiers and Boolean attributes and extend Justicia to verify fairness metrics for more practical settings of decision trees, linear classifiers, and continuous attributes.

Extending to Decision Trees and Linear Classifiers. In the SSAT approach of Section 3, we assume that the classifier \mathcal{M} is represented as a CNF formula. In the literature of interpretable machine learning, several studies have been conducted for learning CNF classifiers in the supervised learning setting, which include but are not limited to the work of (Angelino et al., 2017; Malioutov and Meel, 2018; Ghosh and Meel, 2019). Additionally, we extend Justicia beyond CNF classifiers to decision trees and linear classifiers², which are widely used in the fairness studies (Zemel et al., 2013; Raff et al., 2018; Zhang and Ntoutsis, 2019).

Extending to Continuous Attributes. In practical problems, attributes are generally real-valued or categorical. But classifiers which are already represented using CNF are usually trained on a Boolean abstraction of the input attributes. In order to perform this Boolean abstraction, each categorical attribute is one-hot encoded and each real-valued attribute is discretised into a set of Boolean attributes (Lakkaraju et al., 2019; Ghosh et al., 2020). Detailed design choices are deferred to Appendix B.

4. Empirical Performance Analysis

In this section, we discuss the empirical studies to evaluate the performance of Justicia in verifying different fairness metrics. We first discuss the experimental setup and the objective of the experiments and then evaluate the experimental results.

2. Linear classifiers can be encoded to CNF using pseudo-Boolean encoding (Roussel and Manquinho, 2009).

Table 2: Scalability of different verifiers in terms of execution time (in seconds). DT and LR refer to decision tree and logistic regression respectively. ‘—’ refers to timeout.

Dataset	Ricci		Titanic		COMPAS		Adult	
	DT	LR	DT	LR	DT	LR	DT	LR
Justicia	0.1	0.2	0.1	0.9	0.1	0.2	0.2	1.0
FairSquare	4.8	—	16.0	—	36.9	—	—	—
VeriFair	5.3	2.2	1.2	0.8	15.9	11.3	295.6	61.1

Table 3: Verification of different fairness enhancing algorithms for multiple datasets and classifiers using Justicia. Numbers in bold refer to fairness improvement compared against the unprocessed (orig.) dataset. RW and OP refer to reweighing and optimized-preprocessing algorithm respectively. Results for German dataset is deferred to the Appendix

Classifier	Dataset →	Adult						COMPAS					
	Protected →	Race			Sex			Race			Sex		
	Algorithm →	orig.	RW	OP	orig.	RW	OP	orig.	RW	OP	orig.	RW	OP
Logistic regression	Disparte impact	0.23	0.85	0.59	0.03	0.61	0.62	0.34	0.36	0.47	0.48	0.80	0.74
	Stat. parity	0.09	0.01	0.05	0.16	0.04	0.03	0.39	0.33	0.21	0.23	0.09	0.10
	Equalized odds	0.13	0.03	0.10	0.30	0.02	0.06	0.38	0.33	0.18	0.17	0.19	0.07
Decision tree	Disparte impact	0.82	0.60	0.67	0.00	0.73	0.95	0.61	0.58	0.57	0.94	0.78	0.63
	Stat. parity	0.02	0.05	0.04	0.14	0.05	0.01	0.18	0.17	0.17	0.02	0.09	0.18
	Equalized odds	0.07	0.05	0.03	0.47	0.03	0.04	0.17	0.16	0.16	0.07	0.05	0.16

4.1 Experimental Setup

We have implemented a prototype of Justicia in Python (version 3.7.3). The core computation of Justicia relies on solving SSAT formulas using an off-the-shelf SSAT solver. To this end, we employ the state of the art RE-SSAT solver of (Lee et al., 2017) and the ER-SSAT solver of (Lee et al., 2018). Both solvers output the exact satisfying probability of the SSAT formula.

For comparative evaluation of Justicia, we have experimented with two state-of-the-art distributional verifiers FairSquare and VeriFair, and also a sample-based fairness measuring tool: AIF360. In the experiments, we have studied three type of classifiers: CNF learner, decision trees and logistic regression classifier. Decision tree and logistic regression are implemented using scikit-learn module of Python (Pedregosa et al., 2011) and we use the MaxSAT-based CNF learner IMLI of (Ghosh and Meel, 2019). We have used the PySAT library (Ignatiev et al., 2018) for encoding the decision function of the logistic regression classifier into a CNF formula. We have also verified two fairness-enhancing algorithms: reweighing algorithm (Kamiran and Calders, 2012) and the optimized pre-processing algorithm (Calmon et al., 2017). We have experimented on multiple datasets containing multiple protected attributes: the UCI Adult and German-credit dataset (Dua

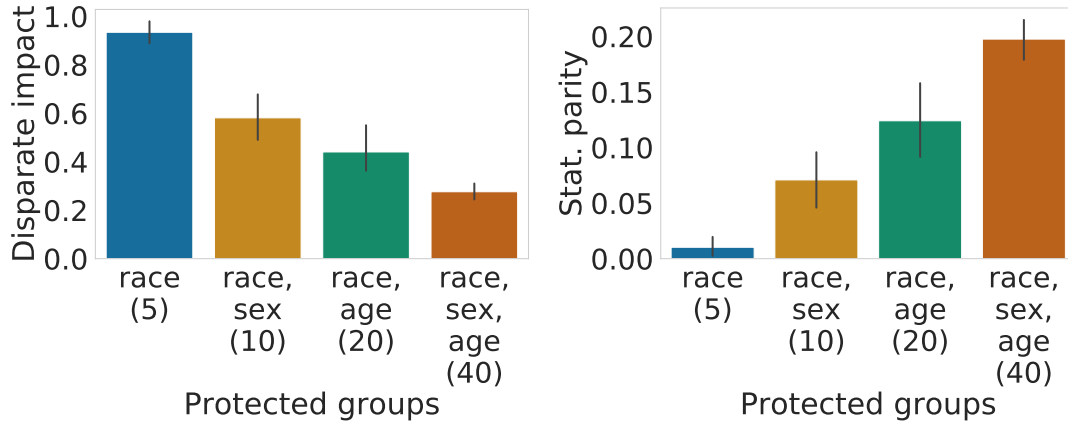


Figure 2: Fairness metrics measured by Justicia for different protected groups in the Adult dataset. The number within parenthesis in the xticks denotes total compound groups.

and Graff, 2017), ProPublicas COMPAS recidivism dataset (Angwin et al., 2016), Ricci dataset (McGinley, 2010), and Titanic dataset³.

Our empirical studies have the following objectives:

1. How accurate and scalable Justicia is with respect to existing fairness verifiers, FairSquare and VeriFair?
2. Can Justicia verify the effectiveness of different fairness-enhancing algorithms on different datasets?
3. Can Justicia verify fairness in the presence of compound sensitive groups?
4. How robust is Justicia in comparison to sample-based tools like AIF360 for varying sample sizes?

Our experimental studies validate that Justicia is more accurate and scalable than the state-of-the-art verifiers FairSquare and VeriFair. Justicia is able to verify the effectiveness of different fairness-enhancing algorithms for multiple fairness metrics, and datasets. Justicia achieves scalable performance in the presence of compound sensitive groups that the existing verifiers cannot handle. Finally, Justicia is more robust than the sample-based tools such as AIF360.

4.2 Experimental Analysis

Accuracy: Less Than 1%-error. In order to assess the accuracy of different verifiers, we have considered the decision tree in Figure 1 for which the fairness metrics are analytically computable. In Table 1, we show the computed fairness metrics by Justicia, FairSquare, VeriFair, and AIF360. We observe that Justicia and AIF360 yield more accurate estimates of DI and SP compared against the ground truth with less than 1% error. FairSquare

3. <https://www.kaggle.com/c/titanic>

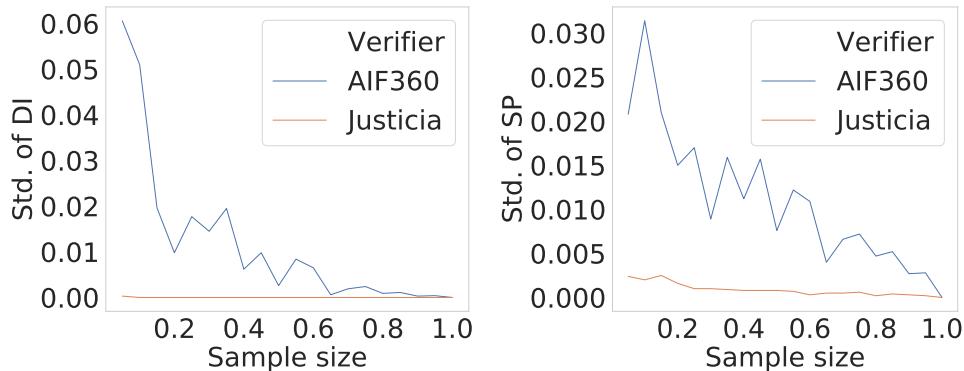


Figure 3: Standard deviation in estimation of disparate impact (DI) and stat. parity (SP) for different sample sizes. *Justicia* is more robust with variation of sample size than AIF360.

and VeriFair estimate the disparate impact to be 0.99 and thus, being unable to verify the fairness violation. Thus, *Justicia* is significantly accurate than the existing formal verifiers: FairSquare and VeriFair.

Scalability: 1 to 3 Magnitude Speed-up. We have tested the scalability of *Justicia*, FairSquare, and VeriFair on practical benchmarks with a timeout of 900 seconds and reported the execution time of these verifiers on decision tree and logistic regression in Table 2. We observe that *Justicia* shows impressive scalability than the competing verifiers. Particularly, *Justicia* is 1 to 2 magnitude faster than FairSquare and 1 to 3 magnitude faster than VeriFair. Additionally, FairSquare times out in most benchmarks. Thus, *Justicia* is not only accurate but also scalable than the existing verifiers.

Verification: Detecting Compounded Discrimination in Protected Groups. We have tested *Justicia* for datasets consisting of multiple protected attributes and reported the results in Figure 2. *Justicia* operates on datasets with even 40 compound protected groups and can potentially scale more than that while the state-of-the-art fairness verifiers (e.g., FairSquare and VeriFair) consider a single protected attribute. Thus, *Justicia* removes an important limitation in practical fairness verification. Additionally, we observe in most datasets the disparate impact decreases and thus, discrimination increases as more compound protected groups are considered. For instance, when we increase the total groups from 5 to 40 in the Adult dataset, disparate impact decreases from around 0.9 to 0.3, thereby detecting higher discrimination. Thus, *Justicia* detects that the marginalized individuals of a specific type (e.g., ‘race’) are even more discriminated and marginalized when they also belong to a marginalized group of another type (e.g., ‘sex’).

Verification: Fairness of Algorithms on Datasets. We have experimented with two fairness-enhancing algorithms: the reweighing (RW) algorithm and the optimized-preprocessing (OP) algorithm. Both of them pre-process to remove statistical bias from the dataset. We study the effectiveness of these algorithms using *Justicia* on three datasets each with two different protected attributes. In Table 3, we report different fairness metrics on logistic regression and decision tree. We observe that *Justicia* verifies fairness improvement as the bias mitigating algorithms are applied. For example, for the Adult dataset with ‘race’

as the protected attribute, disparate impact increases from 0.23 to 0.85 for applying the reweighing algorithm on logistic regression classifier. In addition, statistical parity decreases from 0.09 to 0.01, and equalized odds decreases from 0.13 to 0.03, thereby showing the effectiveness of reweighing algorithm in all three fairness metrics. **Justicia** also finds instances where the fairness algorithms fail, specially when considering the decision tree classifier. Thus, **Justicia** enables verification of different fairness enhancing algorithms in literature.

Robustness: Stability to Sample Size. We have compared the robustness of **Justicia** with AIF360 by varying the sample-size and reporting the standard deviation of different fairness metrics. In Figure 3, AIF360 shows higher standard deviation for lower sample-size and the value decreases as the sample-size increases. In contrast, **Justicia** shows significantly lower ($\sim 10\times$ to $100\times$) standard deviation for different sample-sizes. The reason is that AIF360 empirically measures on a fixed test dataset whereas **Justicia** provides estimates over the data generating distribution. Thus, **Justicia** is more robust than the sample-based verifier AIF360.

5. Discussion and Future Work

Though formal verification of different fairness metrics of an ML algorithm for different datasets is an important question, existing verifiers are not scalable, accurate, and extendable to non-Boolean attributes. We propose a stochastic SAT-based approach, **Justicia**, that formally verifies independence and separation metrics of fairness for different classifiers and distributions for compound protected groups. Experimental evaluations demonstrate that **Justicia** achieves *higher accuracy* and *scalability* in comparison to the state-of-the-art verifiers, FairSquare and VeriFair, while yielding *higher robustness* than the sample-based tools, such as AIF360.

Our work opens up several new directions of research. One direction is to develop SSAT models and verifiers for popular classifiers like Deep networks and SVMs. Other direction is to develop SSAT solvers that can accommodate continuous variables and conditional probabilities by design.

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Appendix A. Proofs of Theoretical Results

Lemma 1 *Solving the ER-SSAT and RE-SSAT problems are NP^{PP} hard (Littman et al., 2001).*

Proof [Proof of Lemma 1] The decision version of ER-SSAT problem is

$$\Phi := \exists a_1, \dots, \exists a_n, \forall^{p_{x_1}} x_1, \dots, \forall^{p_{x_m}} x_m. \Pr[\phi_{\hat{y}}] \geq t,$$

where t is a threshold in $[0, 1]$. It is exactly the an E-MAJSAT (or threshold SAT) problem which is NP^{PP} hard (Littman et al., 2001). If there's no random variable and $t = 1$, ER-SSAT reduces to a SAT problem, which is NP-hard. If there's no existential variable, ER-SSAT reduces to a MAJSAT problem, which is PP-hard. Similar arguments also hold for RE-SSAT problem. \blacksquare

Lemma 2 *Given Eq. (3) and (4), $\Pr[\Phi_{\text{UR}}] = 1 - \Pr[\Phi'_{\text{ER}}]$.*

Proof [Proof of Lemma 2] Both Φ_{UR} and Φ'_{ER} have random quantified variables in the identical order in the prefix. According to the definition of SSAT formulas,

$$\Pr[\Phi_{\text{UR}}] = \min_{a_1, \dots, a_n} \Pr[\phi_{\hat{Y}}] \text{ and } \Pr[\Phi'_{\text{ER}}] = \max_{a_1, \dots, a_n} \Pr[\neg\phi_{\hat{Y}}].$$

We can show the following duality between ER-SSAT and UR-SSAT,

$$\begin{aligned} \Pr[\Phi'_{\text{ER}}] &= \max_{a_1, \dots, a_n} \Pr[\neg\phi_{\hat{Y}}] \\ &= \min_{a_1, \dots, a_n} (1 - \Pr[\phi_{\hat{Y}}]) \\ &= 1 - \min_{a_1, \dots, a_n} \Pr[\phi_{\hat{Y}}] \\ &= 1 - \Pr[\Phi_{\text{UR}}]. \end{aligned}$$

\blacksquare

Lemma 3 *Let $\Phi_{\mathbf{a}}$ be the RE-SSAT formula for computing the PPV of the compound protected group $\mathbf{a} \in A$. If Φ_{ER} is the ER-SSAT formula for learning the most favored group and Φ_{UR} is the UR-SSAT formula for learning the least favored group, then $\max_{\mathbf{a}} \Pr[\Phi_{\mathbf{a}}] = \Pr[\Phi_{\text{ER}}]$ and $\min_{\mathbf{a}} \Pr[\Phi_{\mathbf{a}}] = \Pr[\Phi_{\text{UR}}]$.*

Proof [Proof of Lemma 3] It is trivial that the PPV of most favored group \mathbf{a}_{fav} is the maximum PPV of all compound groups $\mathbf{a} \in A$. Similarly, the PPV of the least favored group $\mathbf{a}_{\text{unfav}}$ is the minimum PPV of all compound groups $\mathbf{a} \in A$.

By construction of the SSAT formulas, the PPV of \mathbf{a}_{fav} and $\mathbf{a}_{\text{unfav}}$ are $\Pr[\Phi_{\text{ER}}]$ and $\Pr[\Phi_{\text{UR}}]$ respectively. Since $\Pr[\Phi_{\mathbf{a}}]$ is the PPV of the compound group \mathbf{a} ,

$$\max_{\mathbf{a}} \Pr[\Phi_{\mathbf{a}}] = \Pr[\Phi_{\text{ER}}] \text{ and } \min_{\mathbf{a}} \Pr[\Phi_{\mathbf{a}}] = \Pr[\Phi_{\text{UR}}].$$

\blacksquare

Theorem 4 For an ER-SSAT problem, the sample complexity is given by

$$k = O\left((n + \ln(1/\delta)) \frac{\ln m}{\ln \epsilon_0}\right),$$

where $\frac{\hat{p}}{p} \leq \epsilon_0$ with probability $1 - \delta$ such that $\epsilon_0 \geq 1$.

Corollary 5 If k samples are considered from the data-generating distribution in Justicia such that

$$k = O\left((n + \ln(1/\delta)) \frac{\ln m}{\ln \epsilon_0}\right),$$

the estimated disparate impact \hat{DI} and statistical parity \hat{SP} satisfy, with probability $1 - \delta$, $\hat{DI} \leq \epsilon_0 DI$, and $\hat{SP} \leq \epsilon_0 SP$.

Proof [Proof of Corollary 5] By Theorem 4, we get that for k samples obtained from the data generating distribution, where

$$k \geq (n + \ln(1/\delta)) \frac{\ln m}{\ln \epsilon_0},$$

the estimated probability of satisfaction for the most and least favoured groups \hat{p}_{max} and \hat{p}_{min} satisfies

$$\hat{p}_{max} \leq \epsilon_0 \max_{\mathbf{a}} \Pr[\Phi_{\mathbf{a}}] \text{ and } \hat{p}_{min} \leq \epsilon_0 \min_{\mathbf{a}} \Pr[\Phi_{\mathbf{a}}].$$

with probability $1 - \delta$. Thus, the estimated value of disparate impact will satisfy

$$\hat{DI} \triangleq \frac{\hat{p}_{max}}{\hat{p}_{min}} \leq \epsilon_0 \frac{p_{max}}{p_{min}} \leq \epsilon_0 DI,$$

and statistical parity will satisfy

$$\hat{SP} \triangleq |\hat{p}_{max} - \hat{p}_{min}| \leq \epsilon_0 |p_{max} - p_{min}| \leq \epsilon_0 SP,$$

with probability $1 - \delta$. ■

Appendix B. Practical Extensions and Design Choices

In this section, we relax assumptions of Boolean classifiers and Boolean attributes and extend Justicia to verify fairness metrics in a more practical setting. We first discuss the input classifiers of Justicia in the following.

B.1 Beyond CNF Classifiers.

In the presented SSAT approach for verifying fairness, we assume the classifier \hat{Y} to be represented as a CNF formula. In the literature of interpretable machine learning, several studies have been conducted for learning CNF classifiers in the supervised learning setting, which include but are not limited to the work of (Angelino et al., 2017; Malioutov and Meel, 2018; Ghosh and Meel, 2019; Yu et al., 2020). However, Justicia can be extended beyond CNF classifiers, in particular to decision trees and linear classifiers that are widely adopted in the ML fairness studies (Zemel et al., 2013; Zafar et al., 2017; Xu et al., 2019; Zhang and Ntoutsis, 2019; Raff et al., 2018; Friedler et al., 2019).

Encoding Decision Trees as CNF. Existing rule-based classifiers, for example, binary decision trees can be trivially encoded as CNF formulas. In the binary decision tree, each node in the tree is a literal. A path from the root to the leaf is a conjunction of literals (hence, a path is a clause) and the tree itself is a disjunction of all paths (or clauses). In order to derive a CNF representation ϕ of the decision tree, we first construct a DNF by considering all paths terminating at leaves with negative class label ($\hat{y} = 0$) and then complement it to a CNF using De Morgan’s rule. Therefore, for any input that is classified positive by the decision tree satisfies ϕ and vice versa. In `Justicia_learn` for learning the least favored group, we can construct a negated CNF classifier in Eq. 4 by only including paths terminating on positive labeled leaves.

Encoding Linear Classifiers as CNF. Linear classifiers on Boolean attributes can be encoded into CNF formulas using pseudo-Boolean encoding (Philipp and Steinke, 2015). We consider a linear classifier $W \cdot X + b \geq 0$ on Boolean attributes X with weights $W \in \mathcal{R}^{|X|}$ and bias $b \in \mathcal{R}$. We first normalize W and b in $[-1, 1]$ and then round to integers so that the decision boundary becomes a pseudo-Boolean constraint, e.g., *at-least* k constraint. We then apply pseudo-Boolean constraints to CNF translation to encode the decision boundary to CNF. This encoding usually introduces additional Boolean variables and results in large CNF. In order to generate a smaller CNF, we can apply thresholding techniques on the weights W to consider attributes with higher weights only. For instance, if the weight $|w_i| \leq \lambda$ for a threshold λ and $w_i \in W$, we can set $w_i = 0$. Thus, lower weighted (hence less important) attributes do not appear in the encoded CNF. Finally, to construct the negated classifier in the SSAT formula in Eq. 4, we encode $W \cdot X + b < 0$ to CNF using *at-most* k encoding.

In practical problems, attributes are generally real-valued or categorical. We next discuss how `Justicia` can work beyond Boolean attributes.

B.2 Beyond Boolean Attributes.

Classifiers that are already represented in CNF are usually trained on a Boolean abstraction of the input attributes where each categorical attribute is one-hot encoded and each real-valued attribute is discretized into a set of Boolean attributes (Lakkaraju et al., 2019; Ghosh et al., 2020). Thus, `Justicia` can verify CNF classifiers readily.

Decision Trees. In case of binary decision tree classifiers, the input attributes are numerical or categorical, but each attribute is compared against a constant in each internal node of the tree. Hence, we fix a Boolean variable for each internal node where the Boolean assignment to the variable decides one of the two branches to choose from the current node.

Linear Classifiers. Linear classifiers are generally trained on numerical attributes where we apply following discretization. Consider a numerical attribute x where w is its weight. We want to discretize x to a set \mathbf{B} of Boolean attributes and recalculate the weights of the variables in \mathbf{B} from w . For discretization, we simply consider interval-based approach where for each interval (or bin) in the continuous space of x , we consider a Boolean variable $b_i \in \mathbf{B}$ such that b_i is assigned \top (or 1) when the attribute-value of x lies within the interval and b_i is assigned \perp (or 0) otherwise. Let μ_i be the mean of the interval where b_i can be \top .

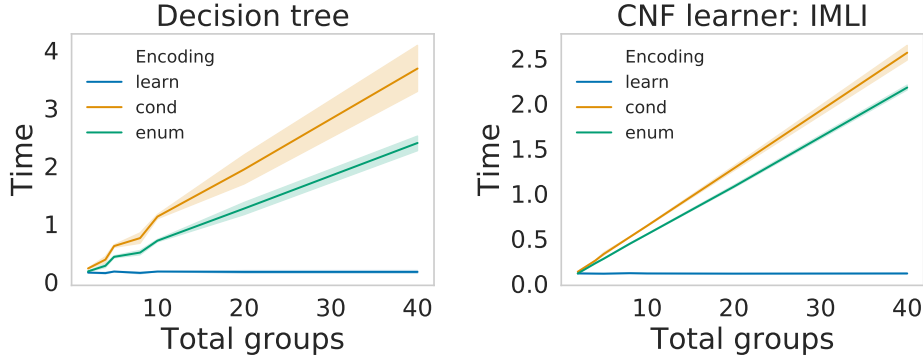


Figure 4: Runtime comparison of different encodings for varying total protected groups in the Adult dataset

We then fix the revised weight of b_i to be $\mu_i \cdot w$. We can show trivially that if we consider infinite number of intervals, $x \approx \sum_i \mu_i b_i$.

Appendix C. Additional Experimental Details

C.1 Experimental Setup

Since both *Justicia* and *FairSquare* take a probability distribution of the attributes as input, we perform five-fold cross validation, use the train set for learning the classifier, compute distribution on the test set and finally verify fairness metrics such as disparate impact and statistical parity difference on the distribution.

C.2 Comparative Evaluation of Two Encodings

While both *Justicia_enum* and *Justicia_learn* have the same output, the *Justicia_learn* encoding improves exponentially in runtime than *Justicia_enum* on both decision tree and Boolean CNF classifiers as we vary the total compound groups in Figure 4. This analysis justifies that the naïve enumeration-based approach cannot verify large-scale fairness problems containing multiple protected attributes.

Table 4: Verification of different fairness enhancing algorithms for multiple datasets and classifiers using *Justicia*. Numbers in bold refer to fairness improvement compared against the unprocessed (orig.) dataset. RW and OP refer to reweighing and optimized-preprocessing algorithm respectively.

Classifier	Dataset →	German					
	Protected →	Age			Sex		
	Algorithm →	orig.	RW	OP	orig.	RW	OP
Logistic regression	Disparte impact	0.00	0.00	0.31	0.27	0.46	0.17
	Stat. parity	0.45	0.03	0.12	0.03	0.02	0.07
	Equalized odds	0.65	0.04	0.14	0.10	0.08	0.13
Decision tree	Disparte impact	0.00	0.56	0.12	0.35	0.37	0.38
	Stat. parity	0.35	0.02	0.22	0.05	0.10	0.11
	Equalized odds	0.36	0.05	0.28	0.06	0.16	0.17